

Simple MCMC under SIR

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We describe how to set-up and run a simple Metropolis-Hastings-based Markov chain Monte Carlo (MCMC) sampler under the susceptible-infected-removed (SIR) model.

```
library(MultiBD)
```

This example uses the Eyam data that consist the population counts of susceptible, infected and removed individuals across several time points.

```
data(Eyam)
Eyam
```

```
##   time    S  I   R
## 1  0.0 254  7   0
## 2  0.5 235 14  12
## 3  1.0 201 22  38
## 4  1.5 153 29  79
## 5  2.0 121 20 120
## 6  2.5 110  8 143
## 7  3.0  97  8 156
## 8  4.0  83  0 178
```

The log likelihood function is the sum of the log of the transition probabilities between two consecutive observations. Note that, we will use $(\log \alpha, \log \beta)$ as parameters instead of (α, β) . The rows and columns of the transition probability matrix returned by `dbd_prob()` correspond to possible values of S (from a to a_0) and I (from 0 to B) respectively.

```
loglik_sir <- function(param, data) {
  alpha <- exp(param[1]) # Rates must be non-negative
  beta  <- exp(param[2])

  # Set-up SIR model
  drates1 <- function(a, b) { 0 }
  brates2 <- function(a, b) { 0 }
  drates2 <- function(a, b) { alpha * b }
  trans12 <- function(a, b) { beta * a * b }

  sum(sapply(1:(nrow(data) - 1), # Sum across all time steps k
    function(k) {
      log(
        dbd_prob( # Compute the transition probability matrix
          t = data$time[k + 1] - data$time[k], # Time increment
          a0 = data$S[k], b0 = data$I[k],      # From: S(t_k), I(t_k)
          drates1, brates2, drates2, trans12,
          a = data$S[k + 1], B = data$S[k] + data$I[k] - data$S[k + 1],
          computeMode = 4, nblocks = 80        # Compute using 4 threads
        )
      )
    }
  ))
}
```

```

    ) [1, data$I[k + 1] + 1] # To: S(t_(k+1)), I(t_(k+1))
  )
})))
}

```

Here, we choose $\text{Normal}(0, 100^2)$ as the prior for both $\log \alpha$ and $\log \beta$.

```

logprior <- function(param) {
  log_alpha <- param[1]
  log_beta <- param[2]

  dnorm(log_alpha, mean = 0, sd = 100, log = TRUE) +
  dnorm(log_beta, mean = 0, sd = 100, log = TRUE)
}

```

We will use the random walk Metropolis algorithm implemented in the function *MCMCmetrop1R()* (**MCMCpack** package) to explore the posterior distribution. So, we first need to install the package and its dependencies.

```

source("http://bioconductor.org/biocLite.R")
biocLite("graph")
biocLite("Rgraphviz")
install.packages("MCMCpack", repos = 'http://cran.us.r-project.org')
library(MCMCpack)

```

The starting point of our Markov chain is the estimated value of (α, β) from Raggett (1982).

```

alpha0 <- 3.39
beta0 <- 0.0212

```

We discard the first 200 iterations and keep the next 1000 iterations of the chain.

```

post_sample <- MCMCmetrop1R(fun = function(param) { loglik_sir(param, Eyam) + logprior(param) },
  theta.init = log(c(alpha0, beta0)),
  mcmc = 1000, burnin = 200)

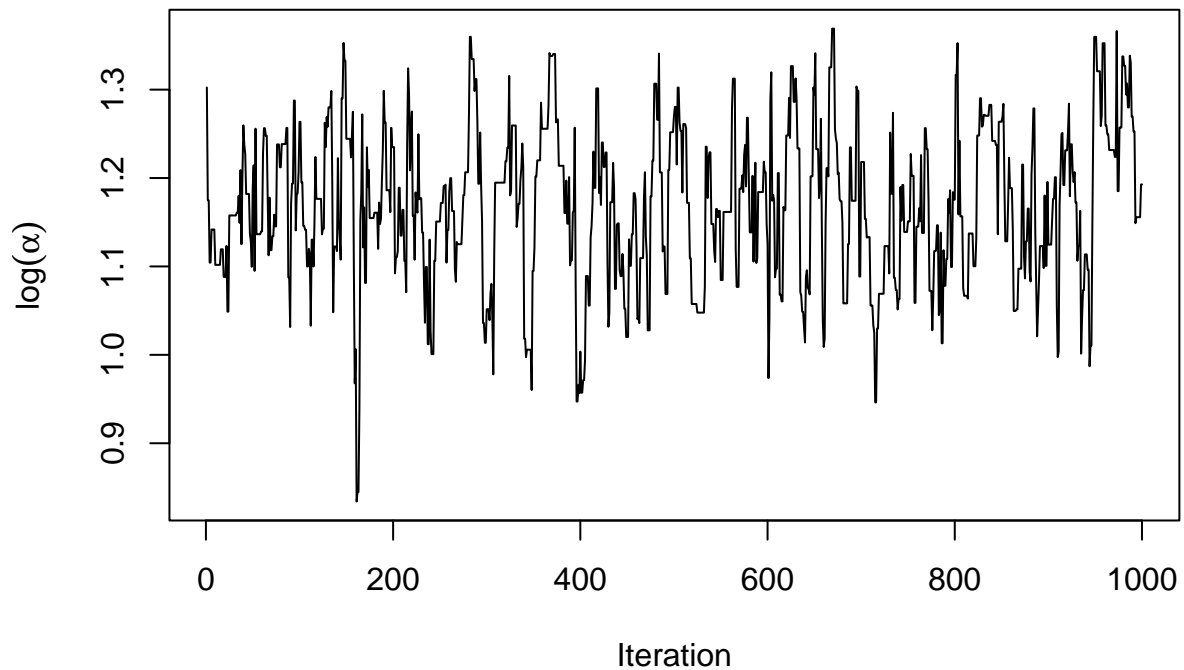
```

The trace plots of both $\log \alpha$ and $\log \beta$ look good.

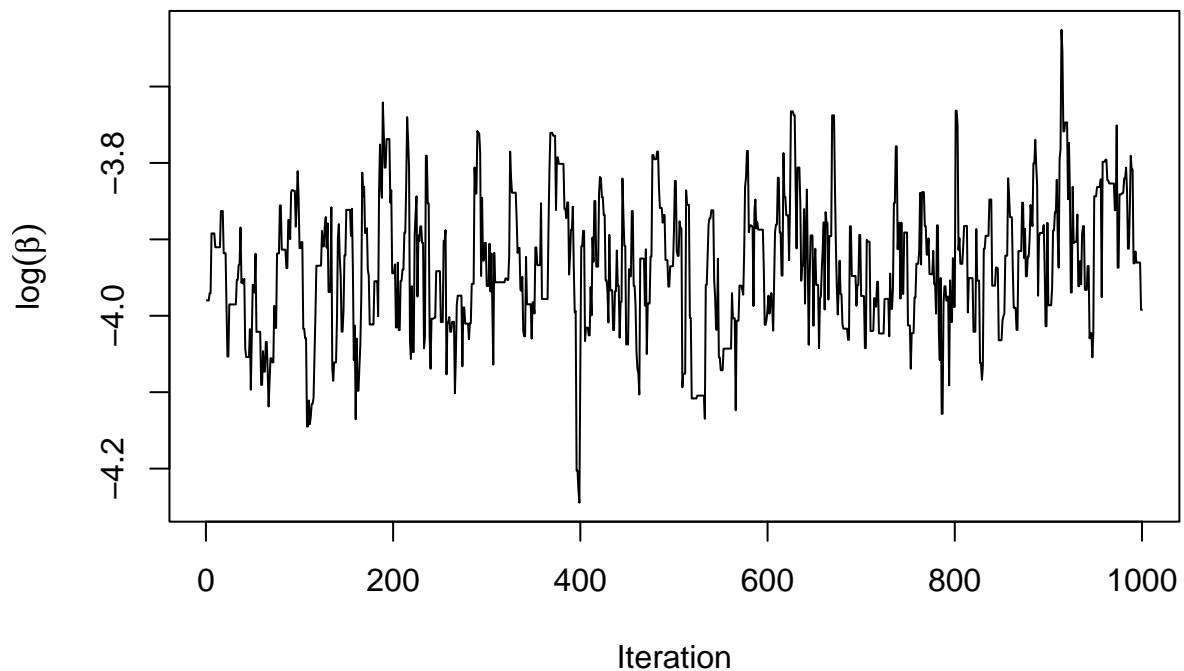
```

plot(as.vector(post_sample[,1]), type = "l", xlab = "Iteration", ylab = expression(log(alpha)))

```

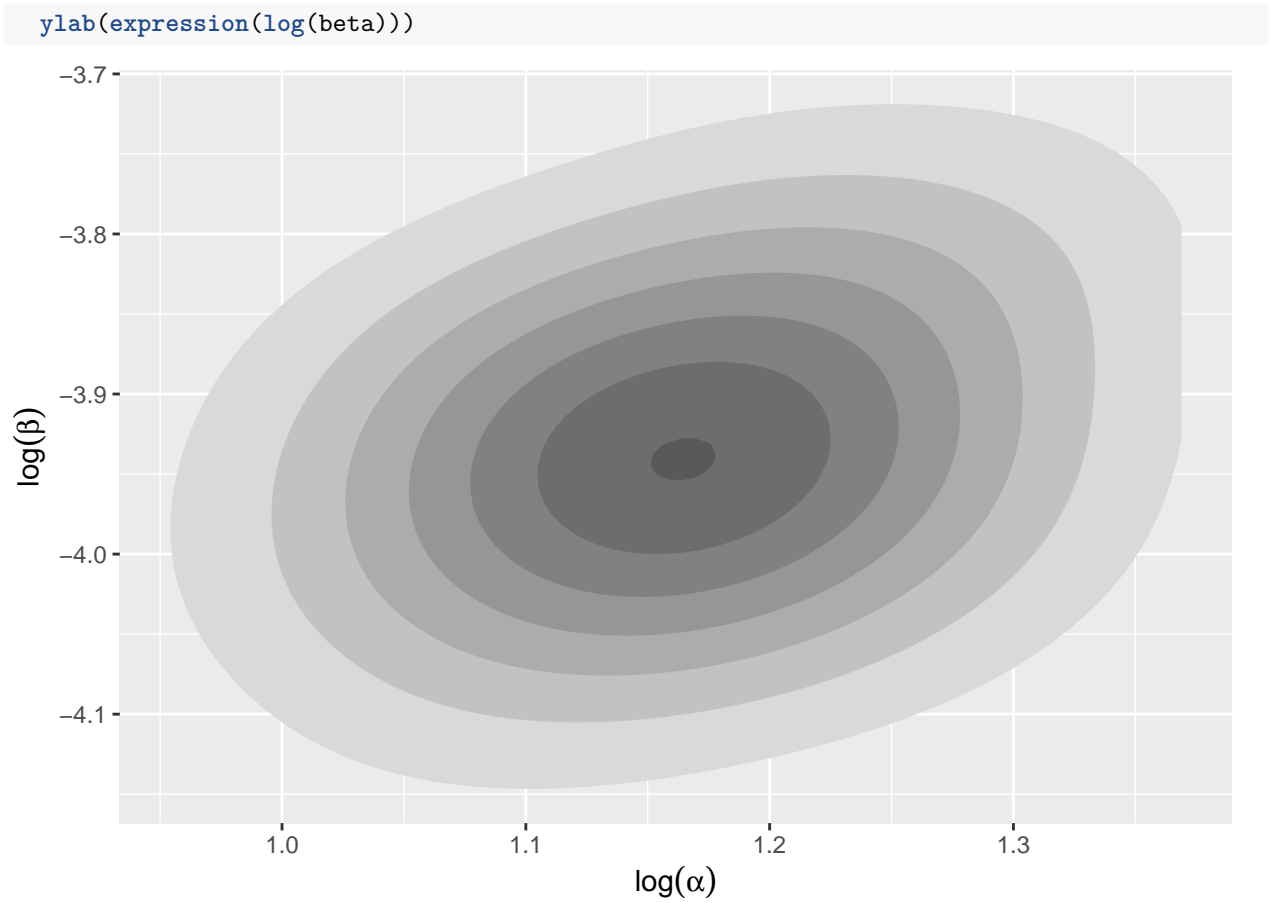


```
plot(as.vector(post_sample[,2]), type = "l", xlab = "Iteration", ylab = expression(log(beta)))
```



We can visualize the joint posterior distribution of $\log \alpha$ and $\log \beta$ using the **ggplot2** package.

```
library(ggplot2)
x = as.vector(post_sample[,1])
y = as.vector(post_sample[,2])
df <- data.frame(x, y)
ggplot(df, aes(x = x, y = y)) +
  stat_density2d(aes(fill = ..level..), geom = "polygon", h = 0.26) +
  scale_fill_gradient(low = "grey85", high = "grey35", guide = FALSE) +
  xlab(expression(log(alpha))) +
```



We can also construct the 95% Bayesian credible intervals for α and β .

```
quantile(exp(post_sample[,1]), probs = c(0.025,0.975))
```

```
##      2.5%      97.5%  
## 2.721921 3.809780
```

```
quantile(exp(post_sample[,2]), probs = c(0.025,0.975))
```

```
##      2.5%      97.5%  
## 0.01649866 0.02327176
```